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Did you really take a hit? Understanding how video games playing affects individuals



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ABSTRACT

This study addresses the important and recurring question of whether playing video games is detrimental to the socio-economic development of a person. It does this by using novel data from the *Taking Part* Survey in England to establish whether games playing is associated with particular socio-economic characteristics and/or other forms of cultural participation. The results do not indicate any obviously negative effects of video games playing: rather, those who play are typically better educated and no less wealthier, and games players are also more likely than non-games players to participate in other forms of culture, particularly active forms of participation. These findings are reinforced when comparing the characteristics of individuals who did and did not play video games when younger.

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1. Introduction

Video games are regarded as one of the most significant and fastest growing form of mass media in recent years (Lenhart, 2015; Dutton et al., 2013). Industry analysts forecast that global video games revenue will reach \$106.6 billion in 2017 (Newzoo, 2017). Every second European plays video games and half of the players do so at least once a week (ISFE, 2012), while the generation that grew up on video games is approaching mid-life. And yet it is still debated whether playing games is a beneficial or harmful activity for individual development. A growing number of studies make use of survey data on educational and psycho-social attitudes and behaviours to examine this question and deliver mixed results (Przybylski and Mishkin, 2016). Here, we use an economic approach and a different source of data – namely, survey data on cultural participation – to see if we can detect positive or negative socio-economic impacts, including in the long run, from playing video games.

The results that we present do not indicate any obviously negative socio-economic effects of video games playing, but suggest rather the opposite: those who play are typically better educated and no less wealthier.¹ With the gamer's average age of 43.2 years the medium cannot be characterized as “immature” and since females are on the whole more likely to play, the activity cannot be described as male-dominated. Games players are also more likely than non-games players to participate in other forms of culture, especially through active participation (e.g., creating digital video, photography or animation or writing stories or poetry), but passive participation (e.g., attending cultural festivals) is also common. These findings are reinforced when comparing the characteristics of individuals who did and did not play video games when younger. In particular, today we do not observe socio-economic differences between those who played when growing up

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¹ The data do not allow us to explore causal effects, but the correlations are robust and rather informative.

and those who did not. We also find that those who played games when growing up participated in other forms of culture too, and in particular they were more likely to read, paint, attend performing arts and visit heritage sites and libraries.

In this paper we make two main contributions. First, the very rich set of cultural variables in the *Taking Part* survey allows us to understand how video games playing is related to cultural consumption by disentangling the effects of *active* cultural participation (e.g., writing a story or poem or painting) from more *passive* forms of cultural participation (e.g., attending events). Both forms of cultural engagement are often associated with positive externalities on the individual, however the former (active participation) more obviously requires or stimulates creative thinking. It therefore comes close to the priorities of cultural policy making in Europe, which focuses heavily on fostering and stimulating creativity in our societies (e.g., [European Commission, 2017](#)). Second, the paper provides unique insights on video games playing early in life and on how games playing when growing up is associated with games playing behaviour later in life. In particular, the results add to the findings from studies of more traditional forms of cultural consumption that cultural tastes persist.

This paper relates thus to the important microeconomic concept of “rational addiction” ([Stigler and Becker, 1977](#)), which has had great influence in the literature on addictive behaviours. Despite the fact that a large volume of applied research on consumption patterns builds heavily on this theory (e.g., “learning-by-consuming” in [Ateca-Amestoy, 2008](#)), it has rarely been formally tested, since data on early life events is typically not available. That the *Taking Part* survey asks individuals to report on their cultural participation when young as well as today means that it can provide new evidence on these theories.²

Besides the core microeconomic literature, the study relates to video games research, which has been growing rapidly over the last few years. [Borowiecki and Prieto-Rodriguez \(2015\)](#) motivate video games playing from the perspective of cultural economics and investigate video game usage in Spain. [Borowiecki and Prieto-Rodriguez \(2017\)](#) explore different profiles of video games players in Denmark depending on the type of games they play. However, most academic research has focused on the supply side of the video games industry and includes analyses of video games console pricing ([Cox, 2008](#)), network effects and product cycles ([Clements and Ohashi, 2005](#)), and outcomes in competitive markets ([Shankar and Bayus, 2003](#)).

2. Theory

We describe below two strands of literature that are relevant to the study. First, we present relevant concepts from behavioural microeconomic theory, in particular the rational addiction model. Second, we summarize relevant studies on cultural consumption which help to motivate our empirical specifications.

[Stigler and Becker \(1977\)](#) famously showed that addictive consumption behaviours, say listening to music or smoking, were consistent with constant preferences if consumers’ appreciation of that good was a function of accumulated consumption. The model has been extended in various ways. [Becker and Murphy \(1988\)](#) have formalized the two conditions required for two aspects of addiction: reinforcement, which implies that with increased consumption, you want to consume even more and tolerance, when given levels of consumption are less satisfying when past consumption has been greater. [Gruber and Koszegi \(2001\)](#) introduce a model that incorporates time-inconsistent preferences. [Laibson \(2001\)](#) illustrates how preferences may change rapidly from moment to moment and explain why temptations should sometimes be avoided. Of relevance for the type of consumption studied here are also models of intertemporal consumption dynamics, including mechanisms such as habit formation ([Taylor and Houthakker, 2010](#)) or learning-by-consuming dynamics ([Levy-Garboua and Montmarquette, 1996](#)).

While the addictiveness of video games is the subject of fierce popular debate, empirical studies of addiction and video games specifically from an economic viewpoint are lacking. The wider literature on behavioural addiction and video games playing is fraught with definitional and conceptual challenges: even if some game players are shown to experience harmful effects from the amount of time playing video games it does not necessarily follow that they have a behavioural addiction. It is safe to say that the jury still remains out on these issues.³

Turning next to the literatures on cultural consumption, we begin with [Bourdieu’s \(1973, 1984\)](#) influential sociological contribution, in which he observes that ‘high’ cultural consumption is especially prevalent among the elites. This has proven to be a robust prediction, true across countries, performing arts activities and time ([Seaman, 2006](#)). After all, it has been partly through cultural consumption that elites have historically been able to separate themselves from other social classes. Consistent with this, income and educational attainment are invariably shown to be important drivers of cultural consumption. Other research has also explored the extent to which taste for the ‘high’ arts is inherited from wealthy parents who have a taste for culture. Although the role of upbringing cannot be neglected, exposure to the arts at school has also been shown to be a significant driver of students’ long-term cultural participation as well, independently of family influence (e.g., [Kracman, 1996](#)).

² Another unique feature of the data is that it allows us to illuminate the profile of video games creators. However, due to the low number of game creators, the analysis we present is only descriptive.

³ As evidenced, for example, by a recent open letter from 26 academic researchers to the World Health Organization on a proposed new category of Gaming Disorders for the International Classification of Diseases: “Our main concerns are the low quality of the research base... The act of formalizing this disorder, even as a proposal, has negative medical, scientific, public-health, societal, and human rights fallout that should be considered.” ([Aarseth et al., 2016](#)). See also the controversies in [Griffiths et al. \(2016\)](#) and [Petry \(2016\)](#).

Cultural consumption, be it attendance at an art museum or playing video games, can be time consuming. It may therefore be affected by factors that limit the available time of a person, such as having young children and/or older people in the household. Other things equal certain types of employment may also constrain the time that is available (e.g., self-employment) and the importance of this constraint will in turn depend on the opportunity cost of the individual's time e.g. measured by forgone wages (Borowiecki, 2004). Since wages data are reliably collected in relatively few surveys, however, studies on cultural consumption often include proxies such as age (e.g., Ateca-Amestoy, 2008). Age variables provide some account for changes in wages over the lifetime of an individual, but also capture other age-related factors such as health.

There exist also significant gender differences in the consumption of culture. Studies show that women are more likely to attend high-brow cultural events than men. These differences arguably reflect gender differences in socialization (Bihagen and Katz-Gerro, 2000) or patterns in the household division of labour (DiMaggio and Mukhtar, 2004). Interestingly, these differences appear not to have declined over time, at least not in the US since the early 80s (DiMaggio and Mukhtar, 2004).

Another feature of cultural consumption is that it often – though not always – occurs in groups, which enables people to share experiences and exchange opinions. Video games are no different in this regard. It has been shown, for example, that both married and unmarried partners spend a considerable amount of time on joint leisure activities (Kalmijn and Bernasco, 2001).

Another determinant of consumption is ease of access. Consumers who live nearby cultural supply, such as in an urban agglomeration, have a much lower cost of access and hence have been shown to be more likely to consume culture than those who live in rural areas (Brook, 2016; Frateschi and Lazzaro, 2008). There can also exist large differences in cultural supply across regions, which are more difficult to explain, and could be related to historical path-dependencies. The geography of cultural supply is in general highly persistent: in the case of Italy, for example, the location of cultural supply remained remarkably persistent over a period of several centuries (Borowiecki, 2015).

Related to cost of access is also a relatively large empirical literature which explores price elasticities of demand (for a review see Seaman, 2005), however estimation is beset by notorious challenges in measuring the quality of cultural products.

3. Data

The Department of Culture, Media and Sport's (DCMS) *Taking Part* Survey is a continuous face-to-face nationally representative household survey of adults aged 16 and over in England. The survey asks respondents detailed questions about the frequency of, reasons for and barriers to participation and attitudes in a very wide range of areas spanning the arts, museums and galleries, archives, libraries, heritage and sport. The survey also collects information on income, wealth, age, education and other economic and socio-demographic variables, and on a large number of other indicators, including cultural participation when "growing up". A variable measuring cultural consumption of a person when growing up can provide potentially novel and important insights into her behaviour in the past. The growing up questions were introduced with the following linking text: "Now I would like to ask you a few questions about activities you may have done when you were growing up. By 'growing up' I mean when you were aged around 11 to 15." At the suggestion of this paper's authors new questions asking about video games consumption and creation activities were introduced by the DCMS into the survey in April 2014. In this paper, we analyse responses gathered in interviews conducted between April 2014 and March 2015.

Table 1 provides descriptive statistics for those who report to play video games and those who do not. The table gives the frequency of playing for those who play and whether any of the two groups of respondent played video games when growing up and, if so, how often. It shows that one in two video games players today also played games when growing up, compared with almost one in eight who do not play nowadays but who did when younger. Both of the frequency measures in *Taking Part* are ordinal variables and take the following six values (we code with 6 the highest frequency): at least once a day, at least once a week, at least once a month, at least 3–4 times a year, at least once a year, or less often than once a year. For games players, at 5 the mean playing frequency corresponds to at least once a week nowadays, a little higher than 4.85 when growing up. When measured by the mode, average games playing frequency is higher, "at least once a day" in the case of playing today. The summary table also gives information on the age, gender, income, education, type of area of residence and physical assets of respondents (in particular whether they own a mobile phone, car or house). In addition, we present information on respondents' employment and family status. In total, there are 9817 individuals in the *Taking Part* sample over this period, of which 1455 have missing observations for income and 32 have missing observations for age.

The boxplots in Fig. 1 depict the age distribution of those who do and do not play video games today and for those who did and did not play when growing up. The shaded boxes visualise the 25th and 75th percentiles and the vertical lines within the boxes indicate the median. Fig. 1 shows that a large proportion of those who play nowadays are older: the median age of 41.5 implies that half of the population of gamers is at least in their forties, while one-fourth of players are 56 years or above. The figure also shows that those who played when growing up tend to be younger (median age of 31), and that there are in fact very few older people who played when growing up. This no doubt reflects the fact that video games are, in relative terms, still a young phenomenon. Taken together, the plots suggest that those who play video games either now or did so when younger are a much younger group than those who do not play games.

Fig. 2 presents a scatter plot with weighted markers to visualise the frequency of playing nowadays and when growing up, and the relationship between these two variables. To denote the importance of each of the dots, their size is weighted by the relative occurrence. Owing to the disproportionately high number of respondents who do not play today and did not

Table 1
Summary statistics.

	Gamers			Non-gamers			Difference
	Obs	Mean	Std. Dev.	Obs	Mean	Std. Dev.	
Play video games	2333	1	0	7484	0	0	
Frequency of playing	2330	5.01	1.06	7484	0	0	5.01*
Played when growing up	1220	0.50	0.50	3607	0.13	0.34	0.37*
Frequency of playing when growing up	230	4.85	0.92	203	4.54	0.89	0.31*
Age	2330	43.20	16.96	7455	56.58	17.84	-13.39*
Female	2333	0.51	0.50	7484	0.57	0.49	-0.06*
Income	2038	5.36	2.95	6324	5.29	2.88	0.07
Educ1 (unknown)	2333	0.02	0.16	7484	0.05	0.21	-0.02*
Educ2 (trade apprentice)	2333	0.03	0.16	7484	0.05	0.22	-0.02*
Educ3 (GCSE, <5 A*-C)	2333	0.06	0.23	7484	0.05	0.22	0.01
Educ4 (GCSE, 5+ A*-C)	2333	0.18	0.38	7484	0.13	0.34	0.05*
Educ5 (A levels)	2333	0.21	0.41	7484	0.15	0.36	0.05*
Educ6 (higher, below degree)	2333	0.11	0.32	7484	0.10	0.31	0.01
Educ7 (higher, degree)	2333	0.28	0.45	7484	0.25	0.43	0.04*
Urban	2333	0.81	0.40	7484	0.79	0.41	0.02*
Wealth (owns mobile)	2332	0.97	0.17	7476	0.88	0.33	0.09*
Wealth (owns car)	2333	0.82	0.39	7484	0.78	0.42	0.04*
Wealth (owns house)	2333	0.26	0.44	7484	0.43	0.50	-0.17*
Self-employed	2333	0.08	0.27	7484	0.07	0.26	0.01
Employed	2333	0.51	0.50	7484	0.38	0.48	0.14*
Married or in cohabitation	2333	0.56	0.50	7484	0.56	0.50	0.00
Number of children under age 4	2333	0.19	0.48	7484	0.13	0.41	0.07*
Number of adults above age 85	2333	0.00	0.06	7484	0.01	0.12	-0.01*

* $p < 0.05$.

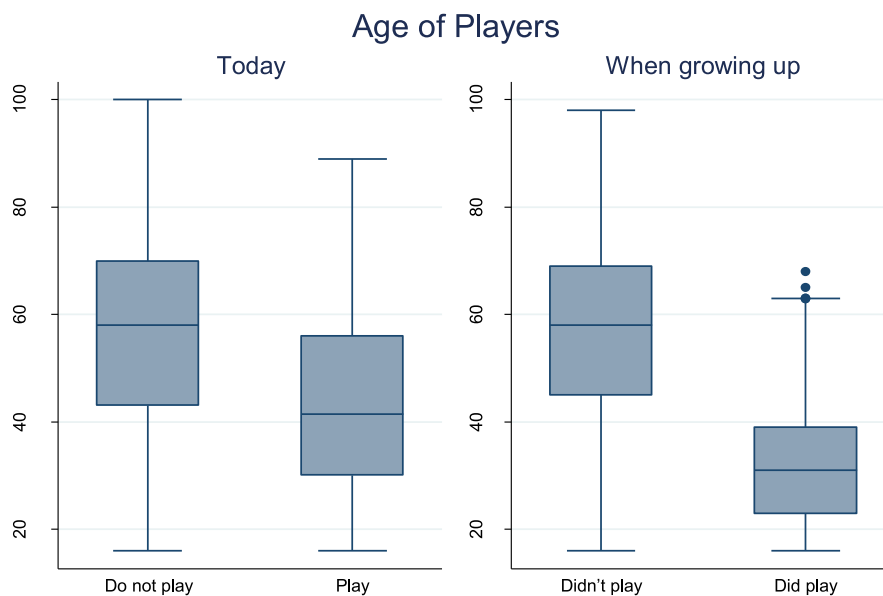


Fig. 1. Age of Players today and of those who played when growing up.

play when growing up (the incidence at the origin of the axes) the weight of these occurrences has been scaled down (and is shown with a shaded marker).

The figure suggests that people either do not play video games at all or, when they do they play with relatively high frequency. Nobody reports playing games when growing up with the frequency of less often than once a year and only 6 (out of 433 players when growing up) report to have played at least once a year (but less than 3–4 times a year), while only 62 respondents (out of 2330 players) nowadays report a frequency in these two lowest frequency categories. Not all who play nowadays played when growing up, while at the same time some of those who played when growing up do not play video games today. Nonetheless, for many respondents we observe persistence in games playing frequency over time: those who played frequently when growing up tend to also play frequently now. Using all the observations, we depict a linear fit line to estimate the correlation between these two variables, which turns out to be a statistically significant 0.27.

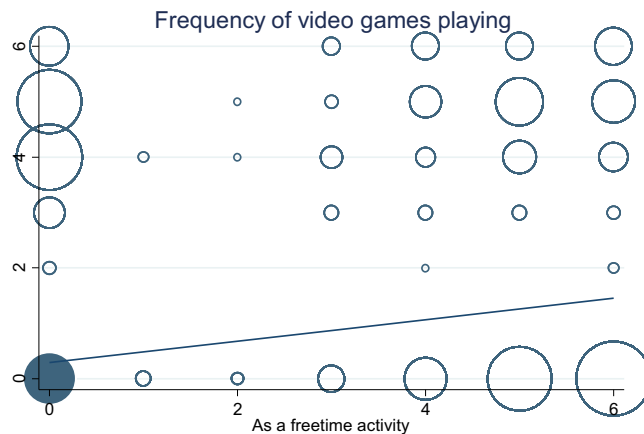


Fig. 2. Frequency of video games playing.

Appendix 1 contains an extended summary table which presents results for the relatively small sample of 107 video games creators. This table presents data on the socio-economic background of creators and their playing probabilities. Interestingly, creators are significantly more likely than others to participate in culture, and in all but one case the difference is statistically significant at the 99% confidence interval.

4. Model

A challenge for modelling participation behaviour in cultural economics using survey data is distinguishing between individuals who report no participation in a cultural activity because they would never participate from those who do in fact have some probability of participating even though they have not done so in the sample period. Bagozzi et al. (2014) show that not treating these individuals as distinct groups introduces model biases (technically, the data is 'zero inflated'). One of the first studies within cultural economics that deals with this heterogeneity is Ateca-Amestoy (2008). The author measures theatre participation using count data and estimates a Zero Inflated Negative Binomial Model by characterizing two distinct behavioural groups: never-goers (who never participate) and a sub-population that has a positive probability of attending.

We estimate standard marginal effects for Probit and ordered Probit models, along with zero-inflated ordered Probit (ZIOP) models which explicitly address the challenges of a "zero-inflated" ordered dependent variable. Below we outline briefly the features of the latter model.⁴ The ZIOP model comprises two equations: a Probit selection equation and a standard ordered Probit. This endogenously splits the observations into two groups that exhibit potentially different associations with the explanatory variables. For the case of playing video games, respondents are assigned into one of the two groups: those who never play and are characterized always by zero consumption, and those who potentially play which may either have positive or null consumption in the period surveyed. In other words, individuals who exhibit zero-consumption patterns are potentially attributed to two different populations. Those in the first group have no interest in video games and exhibit perfectly inelastic demand to personal constraints. Those in the second group also report zero consumption, but may consume once their constraints have been relaxed. In our models, we capture the constraints used to "inflate" the participation equation by four different variables that are intuitively associated with time commitments of a person: whether or not they are employed, self-employed, live together with a partner, have children under the age of four and live in a household with a person older than 85 years.⁵

In order to study reliably the determinants of video games playing, the choice of the underlying variables has to be conducted in line with economic theory and should also be consistent with previous research.⁶ We build on the established approach in the literature of using revealed preferences in order to attribute human behaviour to binding budget constraints, time constraints, social constraints and physical constraints and tastes. Formally, we estimate the following equation:

$$VG_i = f(D_i, I_i, HC_i, G_i, P_i, AP_i, PP_i) \quad (1)$$

The dependent variable, VG_i , is an ordinal variable and it indicates the frequency of a person playing video games as a free time activity. The explanatory variables consist of a set of socio-economic background controls, including demographic variables (D_i) – in our case, age and gender, and income (I_i) – which is reflected in a 12-point ordinal scale. The equation

⁴ Technical details on ZIOP models can be found in Harris and Zhao (2007).

⁵ Including instead (or in addition) a variable measuring the number of children above the age of four and below sixteen does not yield a significant coefficient for this variable (this could be a reflection of the likelihood that some parents may play video games with their older children). The baseline results remain though unchanged.

⁶ A comprehensive review of research on cultural participation is presented by Seaman (2006).

also contains a measure of human capital (HC_i), which is represented by an indicator function taking the value one for each category of the highest level of education obtained by the respondent. Within our framework, and in line with previous literature (see, for example, [Ateca-Amestoy, 2008](#)), the human capital variable is further used as a proxy for the ‘cultural capital’ of a person; that is, the personal resource accumulated by past consumption of, or exposure to, cultural experiences, and – education-based proxies aside – very difficult to quantify. Typically, education, income and occupational status are believed to be the three fundamental determinants of cultural participation ([Gray et al., 2017](#)), and the *Taking Part* data allows us to control for all three of these (occupation status is included in the inflation equation).

In general there exist significant differences between urban and rural areas in terms of ease of access to leisure activities. Therefore, we include a geographical control (G_i) to indicate whether the respondent is based in an urban or rural location. Individuals who want to play video games require hardware, which we account for and label as physical capital (P_i). *Taking Part* does not contain questions regarding the ownership of games consoles or personal computers, so we try and approximate the required physical capital with variables measuring whether or not the respondent owns a mobile phone, a car or a house – though these variables will also of course capture an individual’s overall personal wealth.

The consumption of other cultural goods and services could conceivably also be related to video games playing. Previous research suggests this effect is positive ([Borowiecki and Prieto-Rodriguez, 2015](#)), however, *ex ante* the association is ambiguously signed. On the one hand, since consumption of most cultural products is costly, involvement in one activity substitutes for participation in another. On the other, cultural consumption is addictive, and with increased exposure to one cultural product, people may increase demand for complementary cultural products. We incorporate cultural consumption in the model by reflecting two types of cultural involvement: active participation (AP_i) and passive participation (PP_i). Active participation consists of activities, as opposed to events, whereby the individual is actively engaged and often creative involvement is required. These activities include dance (ballet or other types of dance), music (singing, performing or composing), perform (drama, opera, circus), visual arts (painting, drawing, printmaking, sculpture, including street arts), digital arts (photography or video as an artistic activity, or creation of digital art, including animation), art collection (of artworks or crafts), reading (for enjoyment), or writing (stories, plays or poetry).

The nature of passive participation – attendance at events – is arguably different, as these forms of cultural participation do not necessarily require the same degree or perhaps any creative involvement. We equate passive participation with attendance at cinema, visits to events associated with video or electronic art, visual art events (street art or public art display or installation), culturally specific festivals (including circus and carnivals), performing arts events (drama and pantomime, ballet and other types of dance, opera, musicals and concerts, including classical and jazz) and visits to exhibitions (of art, photography, sculpture, craft or books).

Finally, we include in our model a set of control variables for regions to account for unobserved differences in drivers of cultural participation across England, and a set of month controls to account for any unobserved seasonal differences across the 12 months over which the survey data was collected.⁷

5. Results

5.1. Determinants of video games playing as a free time activity

We begin by exploring the determinants of video games playing today and estimate ZIOP models. The results are presented in [Table 2](#), beginning with the inflation equation in column 1 (estimating the probability of ever playing) and followed by the ordered Probit equation in column 2 (estimating the frequency of playing). The basic model with a set of standard determinants is then extended in columns 3 and 4 by the inclusion of various active and passive cultural participation variables, while column 5 presents a standard ordered Probit model for comparison. All models include region and month fixed effects.

The probability of playing (column 1) decreases with age but at a diminishing rate, and is higher for females. The coefficient on income is positive, but statistically insignificant, suggesting that income does not constrain the incidence of playing, and the probability of playing is higher for those who are better educated. The association with urban residence is insignificant. Those who own a mobile phone or a car are more likely to play, whereas there is no statistically significant effect for those who own a house. The identifying equation is estimated properly with the chosen instruments: being employed or self-employed decreases the probability of playing, as does having children under the age of four or living in a household with a person older than 85 years. The frequency estimation (column 2) implies that age again has a negative but diminishing impact on the frequency of playing. Playing frequency is lower for females and for those with higher incomes (the latter perhaps due to access to alternative, more expensive leisure activities or because higher income individuals have a higher opportunity cost of game time). No statistically significant effect is found for educational, type of area of residence, nor measures of physical capital.

Turning to columns 3 and 4, the baseline model is extended by including the set of cultural participation variables. The likelihood ratio *Chi-square* of 269.9 (p -value < 0.001) tells us that the model as a whole is statistically significant, that is,

⁷ One reason for seasonal variation could be weather dependent. For example, if people play more during periods of poor weather (e.g., winter months) and respond to the survey in such a period, their responses may be affected by their temporarily more intense period of playing video games.

Table 2
Determinants of video games playing as a free time activity.

	(1) ZIOP Play video games	(2) ZIOP Play video games	(3) ZIOP Play video games	(4) ZIOP Play video games	(5) Ordered probit Play video games
Age	−0.140*** (0.0218)	−0.0678*** (0.0129)	−0.144*** (0.0221)	−0.0607*** (0.0121)	−0.0125** (0.00509)
Age ² /100	0.0676*** (0.0164)	0.100*** (0.0150)	0.0729*** (0.0166)	0.0866*** (0.0150)	−0.00641 (0.00519)
Female	0.432*** (0.0728)	−0.644*** (0.0616)	0.365*** (0.0829)	−0.646*** (0.0635)	−0.197*** (0.0339)
Income	0.0142 (0.0133)	−0.0297*** (0.0108)	0.0198 (0.0143)	−0.0358*** (0.0108)	−0.0211*** (0.00614)
Educ1 (unknown)	−0.274 (0.172)	0.269 (0.246)	−0.300* (0.171)	0.197 (0.235)	−0.118 (0.0979)
Educ2 (trade apprentice)	−0.138 (0.148)	0.0285 (0.204)	−0.181 (0.151)	−0.0199 (0.204)	−0.222** (0.0948)
Educ3 (GCSE, <5 A*-C)	0.196 (0.148)	−0.0675 (0.149)	0.194 (0.154)	−0.117 (0.148)	0.0935 (0.0774)
Educ4 (GCSE, 5+ A*-C)	0.263** (0.108)	−0.0105 (0.119)	0.239** (0.114)	−0.0692 (0.119)	0.134** (0.0597)
Educ5 (A levels)	0.185* (0.104)	−0.00895 (0.116)	0.147 (0.109)	−0.0949 (0.116)	0.0423 (0.0591)
Educ6 (higher, below degree)	0.269** (0.115)	−0.0545 (0.128)	0.173 (0.124)	−0.136 (0.128)	0.0482 (0.0656)
Educ7 (higher, degree)	0.222** (0.0997)	−0.00851 (0.116)	0.0740 (0.111)	−0.104 (0.118)	−0.0276 (0.0597)
Urban	0.045 (0.0757)	−0.0047 (0.0697)	0.0901 (0.0804)	−0.0377 (0.0706)	0.0155 (0.0398)
Wealth (owns mobile)	0.436*** (0.107)	−0.257 (0.201)	0.426*** (0.109)	−0.273 (0.201)	0.248*** (0.0744)
Wealth (owns car)	0.169** (0.0844)	−0.0320 (0.0675)	0.178** (0.0889)	−0.0431 (0.0674)	0.0209 (0.0424)
Wealth (owns house)	0.0531 (0.0787)	−0.0580 (0.0886)	0.0215 (0.0826)	−0.0278 (0.0895)	0.0372 (0.0412)
<i>Active participation:</i>					
Read			0.162** (0.0722)	0.103* (0.0578)	0.130*** (0.0352)
Write			0.351** (0.156)	−0.139 (0.102)	0.0346 (0.0678)
Create Visual arts			0.152 (0.101)	0.118* (0.0673)	0.116** (0.0454)
Create Digital art			0.261*** (0.100)	0.0675 (0.0732)	0.138*** (0.0459)
Dance			0.162 (0.107)	−0.0915 (0.0692)	−0.0240 (0.0452)
Music			−0.0592 (0.0952)	0.0425 (0.0745)	0.0248 (0.0464)
Perform			0.167 (0.185)	0.116 (0.115)	0.133* (0.0801)
Collect			0.121 (0.0889)	0.0690 (0.0687)	0.118*** (0.0420)
<i>Passive participation:</i>					
Cinema			0.00242 (0.0715)	0.251*** (0.0583)	0.158*** (0.0345)
Visual arts visits			0.0555 (0.0903)	0.0304 (0.0721)	0.0843* (0.0438)
Digital arts visits			0.0600 (0.198)	−0.0433 (0.131)	−0.00347 (0.0869)
Performance arts attendance			−0.0507 (0.0701)	0.0136 (0.0563)	−0.0189 (0.0344)
Festival attendance			0.329*** (0.0960)	−0.127** (0.0634)	0.0376 (0.0402)
Exhibition attendance			−0.124 (0.0816)	−0.0989 (0.0687)	−0.0968** (0.0413)
Self-employed	−0.340*** (0.102)		−0.410*** (0.112)		
Employed	−0.186*** (0.0694)		−0.213*** (0.0771)		

(continued on next page)

Table 2 (continued)

	(1) ZIOP Play video games	(2) ZIOP Play video games	(3) ZIOP Play video games	(4) ZIOP Play video games	(5) Ordered probit Play video games
Married or in cohabitation	0.0582 (0.0539)		0.0603 (0.0592)		
Number of children under age 4	–0.237** (0.0959)		–0.216* (0.111)		
Number of adults above age 85	–0.625** (0.309)		–0.626* (0.330)		
Observations	8344	8344	8344	8344	8344
Month controls	yes	yes	yes	yes	yes
Region controls	yes	yes	yes	yes	yes

Standard errors in parentheses. *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$.

it fits significantly better than a model with no predictors.⁸ We also confirm that the cultural consumption variables as a group are statistically significant (with an emerging likelihood ratio *chi-square* of 106.9 (p -value < 0.001)).

The probability of playing (column 3) is higher for individuals who read, write (stories or poetry), are involved in creating some form of digital art (photo, video or animation), or attend festivals (festival, circus or carnival). The frequency of playing (column 4) increases for those who read, create visual art (paint or street arts), or go to the cinema, but decreases for those who attend festivals. The size and significance of the coefficients on the demographic, income and physical capital variables are broadly unchanged. However, education becomes insignificant. This supports the earlier suggestion that the education variables might proxy for cultural capital as well as formal education.

Two further observations can be made. First, the vast majority of coefficients for cultural participation are positively signed and those that are significant (discussed above) are estimated with a high degree of statistical confidence. Second, most of the significant coefficients relate to forms of active cultural participation – that is, when the individual is creating (e.g., writes stories or poetry, or creates digital art), – whereas for passive participation (e.g., attendance at art events), the coefficients are estimated with lower precision.

The final column reports an ordered Probit model that depicts the correlates of the probability of video games playing. The background variables are estimated to be consistent with the ZIOP models. While the correlation coefficients for some of the cultural activities change, the previous observation that primary active cultural participation is significantly and positively related to playing frequency remains robust.

5.2. Video games playing when growing up

The *Taking Part* Survey contains several questions asking respondents about their cultural participation when growing up, including whether they played “video/computer games on a digital device like a console, a computer, a tablet or a mobile phone”. This question was asked of half of the sample and responses are available for 4827 respondents. Columns 1 and 2 of Table 3 present the marginal effects for a Probit model.⁹

The point estimates are negative for age, indicating that those who played when growing up are younger nowadays. Female are less likely to have played when growing up: the probability of women having played when growing up is 7.9% lower than for men, all else equal. Those who played when growing up tend to have obtained better educational qualifications, but are less likely to own a house.

In the next column we introduce the available measures for cultural participation when growing up. While the estimates for age and female remain broadly unchanged, the association with higher education nowadays disappears. This could once again be an artefact of the complementary association between human capital and cultural capital, which is likely to have accumulated through exposure to various cultural activities when growing up. Thus this finding provides further support to the idea (e.g., *Ateca-Amestoy, 2008*) that education attainment serves as a proxy for cultural capital. The coefficients for cultural consumption when growing up turn mostly positive and statistically significant in five out of eight cases, suggesting that individuals who played video games when growing up, were also significantly more likely to be involved in other cultural activities. In particular, those who played were more likely to read for pleasure, paint or draw, attend performing arts (go to theatre, dance or classical music performances), visit historic sites or libraries. These results are in line not only with the baseline results, but illuminate also a striking persistence in the preferences and behaviours of the respondents. In particular, those who play video games consume more culture, and this relationship was true earlier in life, as it is nowadays.

⁸ The model presented in columns 1 and 2 had a slightly lower likelihood ratio *chi-square* coefficient of 259.4.

⁹ It is not possible to utilize a ZIOP model in this setting due to the lack of adequate variables that could be used in the inflation equation (i.e., nothing is known about the respondents' socio-economic characteristics when growing up).

Table 3
Video games playing when growing up.

	(1) Probit (marginal) Played when growing	(2) Probit (marginal) Played when growing
Age	0.00141 (0.00140)	0.00176 (0.00122)
Age ² /100	−0.0104*** (0.00117)	−0.0100*** (0.00107)
Female	−0.0788*** (0.0126)	−0.0834*** (0.0138)
Income	0.000861 (0.00102)	0.000715 (0.000919)
Educ1 (unknown)	−0.0134 (0.0170)	−0.0175 (0.0129)
Educ2 (trade apprentice)	0.0235 (0.0256)	0.0106 (0.0202)
Educ3 (GCSE, <5 A*-C)	0.0329 (0.0229)	0.0216 (0.0192)
Educ4 (GCSE, 5+ A*-C)	0.0184 (0.0157)	0.00611 (0.0126)
Educ5 (A levels)	0.0303* (0.0170)	0.0104 (0.0131)
Educ6 (higher, below degree)	0.0230 (0.0179)	0.00501 (0.0133)
Educ7 (higher, degree)	0.0290* (0.0155)	0.00385 (0.0118)
Urban	0.00806 (0.00686)	0.00589 (0.00624)
Wealth (owns mobile)	−4.77e−05 (0.0187)	−4.26e−05 (0.0170)
Wealth (owns car)	0.00756 (0.00662)	0.00694 (0.00597)
Wealth (owns house)	−0.0154* (0.00891)	−0.0129 (0.00808)
<i>When growing up:</i>		
Read		0.0114** (0.00549)
Wrote		−0.00111 (0.00552)
Painted		0.0123** (0.00562)
Performed		0.00812 (0.00555)
Attended performing arts		0.0109* (0.00604)
Visited museums		0.00464 (0.00595)
Visited heritage sites		0.0110* (0.00606)
Visited libraries		0.0102* (0.00608)
Observations	3992	3992
Month controls	Yes	yes
Region controls	Yes	yes

Standard errors in parentheses. *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$.

The included socio-economic variables do not obviously indicate any harmful long-term effects of video games playing. The differences in income, educational attainment, place of residence or wealth between those who have and have not played when growing up are mostly statistically insignificant. With the exception of the negative coefficient for house ownership, the variables that turn significant in one of the two models suggest, if anything, the existence of positive correlations. For example, those who have played when growing up have likely obtained nowadays more advanced educational qualifications. These are very general associations and some degree of caution is required in their interpretation, nonetheless, the results confront the stereotypical perception about harmful effects of video games playing.

5.3. Persistence in video games playing

We turn next to the analysis of the effects of video games playing when growing up on playing behaviour today. For this reason, we estimate Probit models with the probability of playing nowadays a function of the various measures of

Table 4
Persistence in video games playing.

	(1) Probit (marginal) Play video games	(2) Probit (marginal) Play video games	(3) Ordered probit Frequency of playing
Played video games when growing up	0.273*** (0.0228)	0.263*** (0.0229)	
Read		0.00698 (0.0173)	
Wrote		0.0363** (0.0170)	
Painted		0.00417 (0.0152)	
Performed		0.0347** (0.0156)	
Attended performing arts		−0.00731 (0.0163)	
Visited museums		−0.0139 (0.0176)	
Visited heritage sites		0.0666*** (0.0172)	
Visited libraries		0.0156 (0.0188)	
Frequency of playing when growing up			0.119*** (0.0231)
Observations	3992	3992	1572
Socio-economic controls	yes	yes	yes
Month controls	yes	yes	yes
Region controls	yes	yes	yes

Standard errors in parentheses. *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$.

cultural participation, including video games playing when growing up. The models contain control variables for individuals' socio-economic backgrounds, regions and month (not reported).

In column 1 of [Table 4](#) we include first the dummy variable indicating whether or not the respondent was playing video games when growing up and present the marginal effects. The point estimate is positive, highly significant and rather large, implying that those who have played when growing up are 27.3% more likely to play nowadays.

Next, the model is extended in column 2 by including indicators of other forms of cultural participation when growing up. We find several positive and significant associations between video games playing today and the newly added variables. In particular, the probability of playing nowadays is higher for those who wrote (stories, poems, plays or music), performed (acted, played musical instrument, sung or danced), and visited historic sites (e.g., heritage buildings, historic parks and gardens and archaeological sites). We note that the estimated effect of playing video games when growing on playing today remains stable and is also the largest in size. This supports the significant and positive association between playing nowadays and when growing up.

Finally, we explore whether playing video games today and when growing up differ depending on the frequency of each of the involvements. However, the data present a challenge in this regard, since the question about the frequency is asked of individuals only in relation to one of their activities when growing up. For individuals who have participated in more than one cultural activity when growing up, the frequency question is asked for only one randomly selected activity. This means that of the 1101 positive occurrences of playing when growing up we obtain frequency data only for 433 cases. To be able to use this in econometric analysis, we need to construct a random sample of an identical weight for the negative cases, that is, those who do not play when growing up. We then repeat this sampling procedure with replacement 50 times in order to obtain the bootstrap estimates of the standard errors. Using the obtained frequency of playing measures, we present in column 3 a model in which we estimate the effect of the frequency of playing when growing up on the frequency of playing today. The point estimates are positive and again highly significant, implying that those who played more frequently when growing up play also more often nowadays.

5.4. Robustness of the results

We begin by assessing whether our results hold when we control for access to the Internet, which could be related to playing video games in at least two ways. First, having internet access increases the attraction of playing, since online gaming (i.e., playing with other people via the Internet) becomes possible.¹⁰ Second, with internet access, greater information

¹⁰ See, for example, [Castronova \(2001\)](#) for an overview of the extent and potential of online video games playing.

Table 5
Robustness tests.

	(1) ZIOP Baseline Play video games	(2) ZIOP Baseline Frequency of playing	(3) ZIOP Internet use Play video games	(4) ZIOP Internet use Frequency of playing	(5) ZIOP Internet access Play video games	(6) ZIOP Internet access Frequency of playing
Age	−0.0607*** (0.0121)	−0.144*** (0.0221)	−0.0553*** (0.0121)	−0.144*** (0.0223)	−0.0552*** (0.0119)	−0.143*** (0.0279)
Age ² /100	0.0866*** (0.0150)	0.0729*** (0.0166)	0.0802*** (0.0146)	0.0780*** (0.0169)	0.0807*** (0.0149)	0.0801*** (0.0204)
Female	−0.646*** (0.0635)	0.365*** (0.0829)	−0.647*** (0.0644)	0.367*** (0.0875)	−0.512*** (0.0624)	0.347*** (0.0880)
Income	−0.0358*** (0.0108)	0.0198 (0.0143)	−0.0391*** (0.0109)	0.0136 (0.0145)	−0.0386*** (0.0105)	−0.000679 (0.0155)
Educ1 (unknown)	0.197 (0.235)	−0.300* (0.171)	0.202 (0.238)	−0.385** (0.175)	0.186 (0.225)	−0.356** (0.178)
Educ2 (trade apprentice)	−0.0199 (0.204)	−0.181 (0.151)	−0.0618 (0.204)	−0.241 (0.158)	−0.0951 (0.195)	−0.220 (0.164)
Educ3 (GCSE, <5 A*-C)	−0.117 (0.148)	0.194 (0.154)	−0.123 (0.149)	0.118 (0.159)	−0.170 (0.140)	0.148 (0.164)
Educ4 (GCSE, 5+ A*-C)	−0.0692 (0.119)	0.239** (0.114)	−0.100 (0.120)	0.132 (0.119)	−0.124 (0.114)	0.142 (0.121)
Educ5 (A levels)	−0.0949 (0.116)	0.147 (0.109)	−0.138 (0.117)	0.00920 (0.115)	−0.138 (0.112)	−0.0472 (0.118)
Educ6 (higher, below degree)	−0.136 (0.128)	0.173 (0.124)	−0.188 (0.130)	0.0268 (0.131)	−0.157 (0.124)	0.0158 (0.132)
Educ7 (higher, degree)	−0.104 (0.118)	0.0740 (0.111)	−0.139 (0.120)	−0.113 (0.116)	−0.115 (0.116)	−0.141 (0.120)
Urban	−0.0377 (0.0706)	0.0901 (0.0804)	−0.0504 (0.0725)	0.088 (0.0829)	−0.028 (0.0689)	0.0629 (0.0861)
Wealth (owns mobile)	−0.273 (0.201)	0.426*** (0.109)	−0.264 (0.199)	0.291** (0.116)	−0.212 (0.185)	0.277** (0.114)
Wealth (owns car)	−0.0431 (0.0674)	0.178** (0.0889)	−0.0617 (0.0682)	0.0859 (0.0938)	−0.0752 (0.0672)	0.0497 (0.0967)
Wealth (owns house)	−0.0278 (0.0895)	0.0215 (0.0826)	0.00487 (0.0932)	−0.0312 (0.0859)	0.0620 (0.0903)	−0.0320 (0.0894)
Self-employed		−0.410*** (0.112)		−0.411*** (0.114)		−0.462*** (0.124)
Employed		−0.213*** (0.0771)		−0.187** (0.0791)		−0.226*** (0.0848)
Married or in cohabitation		0.0603 (0.0592)		0.0527 (0.0614)		−0.0340 (0.0662)
Number of children under age 4		−0.216* (0.111)		−0.247** (0.109)		−0.220* (0.132)
Number of adults above age 85		−0.626* (0.330)		−0.673** (0.337)		−0.631* (0.349)
<i>Active participation:</i>						
Read	0.103* (0.0578)	0.162** (0.0722)	0.0787 (0.0583)	0.107 (0.0746)	0.103* (0.0561)	0.127 (0.0778)
Write	−0.139 (0.102)	0.351** (0.156)	−0.137 (0.103)	0.338** (0.158)	−0.132 (0.100)	0.407** (0.171)
Create Visual arts	0.118* (0.0673)	0.152 (0.101)	0.123* (0.0686)	0.177* (0.104)	0.130** (0.0662)	0.129 (0.108)
Create Digital arts	0.0675 (0.0732)	0.261*** (0.100)	0.0613 (0.0743)	0.207** (0.101)	0.0272 (0.0718)	0.161 (0.108)
Dance	−0.0915 (0.0692)	0.162 (0.107)	−0.113 (0.0701)	0.162 (0.111)	−0.0658 (0.0688)	0.0939 (0.116)
Music	0.0425 (0.0745)	−0.0592 (0.0952)	0.0365 (0.0756)	−0.0878 (0.0963)	0.0479 (0.0733)	−0.0733 (0.102)
Perform	0.116 (0.115)	0.167 (0.185)	0.141 (0.117)	0.161 (0.184)	0.0797 (0.114)	0.182 (0.202)
Collect	0.0690 (0.0687)	0.121 (0.0889)	0.0551 (0.0696)	0.120 (0.0913)	0.0485 (0.0672)	0.134 (0.0961)
<i>Passive participation:</i>						
Cinema	0.251*** (0.0583)	0.00242 (0.0715)	0.238*** (0.0595)	−0.0418 (0.0738)	0.191*** (0.0574)	−0.0436 (0.0760)
Visual arts visits	0.0304 (0.0721)	0.0555 (0.0903)	0.0390 (0.0735)	0.0247 (0.0921)	−0.00382 (0.0708)	0.0618 (0.0977)

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Table 5 (continued)

	(1) ZIOP Baseline Play video games	(2) ZIOP Baseline Frequency of playing	(3) ZIOP Internet use Play video games	(4) ZIOP Internet use Frequency of playing	(5) ZIOP Internet access Play video games	(6) ZIOP Internet access Frequency of playing
Digital arts visits	−0.0433 (0.131)	0.0600 (0.198)	−0.0470 (0.133)	0.0684 (0.197)	−0.204 (0.125)	0.181 (0.232)
Performance arts attendance	0.0136 (0.0563)	−0.0507 (0.0701)	−0.00802 (0.0574)	−0.0503 (0.0725)	−0.000632 (0.0549)	−0.0389 (0.0747)
Festival attendance	−0.127** (0.0634)	0.329*** (0.0960)	−0.138** (0.0646)	0.324*** (0.0989)	−0.127** (0.0627)	0.400*** (0.104)
Exhibition attendance	−0.0989 (0.0687)	−0.124 (0.0816)	−0.118* (0.0701)	−0.138 (0.0840)	−0.0476 (0.0679)	−0.170* (0.0874)
Internet (use)			0.443*** (0.0943)	0.736*** (0.0851)		
Internet (home computer)					−0.0208 (0.0888)	0.719*** (0.0887)
Internet (games console)					0.731*** (0.0738)	1.166*** (0.255)
Internet (mobile)					0.136* (0.0694)	0.424*** (0.0814)
Internet (portable games console)					0.366*** (0.119)	0.369 (0.314)
Observations	8344	8344	8344	8344	8344	8344
Month controls	yes	yes	yes	yes	yes	yes
Region controls	yes	yes	yes	yes	yes	yes

Standard errors in parentheses. *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$.

about the latest video games becomes available – whether via formal advertisement channels or through social networks – and this may increase consumption.

The baseline ZIOP models are re-estimated using additional control variables to account for the respondent's use or type of access to the Internet. Table 5 shows that the baseline results (repeated in columns 1 and 2 of Table 5) remain broadly unchanged if we add dummy variables for whether or not the individual uses the Internet in her spare time (columns 3 and 4), or has accessed the Internet at home using a computer, a games console, a mobile device or a portable games console (columns 5 and 6). As expected, the coefficient estimates for each of these additional control variables are positive and highly significant.

We also explore the existence of any difference between those who play or those who play more often, with regard to nationality, ethnic or religious background, or sexual orientation, but do not find any significant effects of these variables. Nor does inclusion of these variables as additional controls alter the baseline findings.¹¹

6. Conclusions

Playing video games has come of age as a leisure activity, and the generation that grew up on video games is now reaching mid-life. Yet, despite the growing number of studies using survey data on educational and psycho-social attitudes and behaviours to examine the impact of games playing on personal development, a consensus remains lacking. This study has made use of a novel data set – drawn from a national survey of cultural participation in England – to examine the socio-economic characteristics of games players and their wider cultural behaviours. Furthermore, the structure of the *Taking Part* Survey allows us to look at the longer term patterns of video games playing, and in particular its persistence over time.

The results, first of all, do not reveal any obviously negative associations with socio-economic variables. If anything, those who play games are typically better educated, wealthier and are greater consumers of culture. The presented correlations are not proof of causal relationships; nonetheless, the results challenge the stereotypical perception that playing video games has harmful effects on the individual. While there is some evidence that more frequent games playing is associated with lower income, several explanations are consistent with this finding (including the possibility that individuals with lower incomes have a lower opportunity cost of spending time playing video games).

Neither do the results lend any support to the idea that playing video games has longer term detrimental impacts on an individual's socio-economic circumstances. That is, in general, we do not observe any socio-economic differences between those who played when growing up and those who did not. Interestingly, we find that those who played when growing

¹¹ Results are available from the authors on request.

up also participated in other forms of culture, in particular, they were more likely to read, paint, attend performing arts and visit heritage sites or libraries. We also observe a path-dependency in playing video games: those who played when growing up are also more likely to do so nowadays.

Third, although we confirm that video games playing is more popular among younger individuals, with the gamer's average age in *Taking Part* of 43.2 years, the recurring perception that the medium is “immature” is far from being supported. Interestingly, the average gamer is more likely to be female, however among those who play, females do so less often than males.

The study delivers several insights which may have implications for policy makers. The results provide further support for the view that involvement in cultural activities when growing up corresponds with accumulation of cultural capital and has significant path dependencies over the course of life. Cultural policy makers may therefore be right to focus so much on the behaviours of young people. The complementarities we have detected between games playing and other forms of cultural participation may however challenge policy makers' conceptions of what constitutes ‘culture’ from the viewpoint of policy (Bakhshi and Cunningham, 2016).

An important question is whether video games playing makes people more creative. This is a matter we are unfortunately unable to address properly with the data used here. We observe, however, that those who play video games have a stronger tendency to engage in more active forms of cultural activity which is consistent at least with greater creativity. However, it remains an open question whether this finding is a reflection of a creativity-enhancing effect of playing video games, or is rather driven by self-selection of a certain type of person towards both activities – video games playing and active forms of cultural participation.

Appendix

Appendix 1. Video games creators

The *Taking Part* Survey data for the period we consider contains 107 observations for individuals who had created a video game in the previous 12 months. This small number of observations limits econometric analysis, but it is large enough to derive some indicative insights.

Table A1 presents an extended summary statistics. The first three columns compare the survey responses of those who had created a video game with those who play video games (columns 4–6) and those who do not (columns 7–9). These results show that creators are younger than the average player or non-player, typically male, live in urban areas, have above-average income and have obtained higher levels of educational qualification. A bit more than half of video games creators play video games in their free time, which is more than twice the probability for non-creators. Furthermore, some of the creators may regard playing as part of their work as opposed to a free time activity.

Table A2 provides a detailed overview of cultural activities in which creators and non-creators are involved, along with the resulting difference between the two groups (last column). Interestingly, creators are significantly more likely to be participating in culture and in all but one case the difference is statistically significant at the 99% confidence interval.

Table A1

Summary statistics for creators, gamers and non-gamers.

	Creators			Gamers			Non-gamers		
	Obs	Mean	Std. Dev.	Obs	Mean	Std. Dev.	Obs	Mean	Std. Dev.
Play video games	107	0.56	0.50	2333	1	0	7484	0	0
Frequency of playing	105	2.72	2.56	2330	5.01	1.06	7484	0	0
Played when growing up	59	0.56	0.50	1220	0.50	0.50	3607	0.13	0.34
Frequency of playing when growing up	9	5.00	0.71	230	4.85	0.92	203	4.54	0.89
Age	107	40.95	15.47	2330	43.20	16.96	7455	56.58	17.84
Female	107	0.36	0.48	2333	0.51	0.50	7484	0.57	0.49
Income	89	6.55	3.57	2038	5.36	2.95	6324	5.29	2.88
Wealth (owns house)	107	0.20	0.40	2333	0.26	0.44	7484	0.43	0.50
Wealth (owns car)	107	0.79	0.41	2333	0.82	0.39	7484	0.78	0.42
Educ1 (unknown)	107	0.04	0.19	2333	0.02	0.16	7484	0.05	0.21
Educ2 (trade apprentice)	107	0.01	0.10	2333	0.03	0.16	7484	0.05	0.22
Educ3 (GCSE, <5 A*-C)	107	0.04	0.19	2333	0.06	0.23	7484	0.05	0.22
Educ4 (GCSE, 5+ A*-C)	107	0.16	0.37	2333	0.18	0.38	7484	0.13	0.34
Educ5 (A levels)	107	0.17	0.38	2333	0.21	0.41	7484	0.15	0.36
Educ6 (higher, below degree)	107	0.12	0.33	2333	0.11	0.32	7484	0.10	0.31
Educ7 (higher, degree)	107	0.45	0.50	2333	0.28	0.45	7484	0.25	0.43
Urban	107	0.85	0.36	2333	0.81	0.40	7484	0.79	0.41

Creators are identified as those who respond positively to the question: "In the last 12 months, have you created a video or computer game?".

Table A2

The differences between creators and non-creators.

	Creators			Non-creators			Difference
	Obs	Mean	Std. Dev.	Obs	Mean	Std. Dev.	Mean
<i>Active participation:</i>							
Dance	107	0.25	0.44	9817	0.13	0.33	0.13*
Music	107	0.30	0.46	9817	0.13	0.33	0.17*
Perform	107	0.16	0.37	9817	0.03	0.18	0.13*
Create Visual Arts	107	0.25	0.44	9817	0.13	0.33	0.13*
Create Digital Arts	107	0.52	0.50	9817	0.12	0.33	0.40*
Collect	107	0.33	0.47	9817	0.18	0.38	0.15*
Read	107	0.77	0.43	9817	0.66	0.47	0.11*
Write	107	0.25	0.44	9817	0.05	0.22	0.20*
<i>Passive participation:</i>							
Cinema	107	0.73	0.45	9817	0.51	0.50	0.22*
Digital Arts visits	107	0.19	0.39	9817	0.03	0.17	0.16*
Visual arts visits	107	0.37	0.49	9817	0.17	0.38	0.20*
Festival attendance	107	0.22	0.42	9817	0.16	0.37	0.07
Performing arts attendance	107	0.77	0.43	9817	0.57	0.50	0.20*
Exhibition attendance	107	0.43	0.50	9817	0.25	0.43	0.18*

Note: Difference is calculated as the difference between the mean for creators and mean for non-creators. * denotes difference coefficients that are statistically significant at the 1% level.

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